

Identification

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Definition of Identification

Another term that means different things to different people

I will base my discussion on Matzkin's (2007) formal definition of identification but use my own notation and be a bit less formal

This will all be about the **Population** in thinking about identification we will completely ignore sampling issues

We first need to define a data generation process

Data Generating Process

Let me define the data generating process in the following way

$$X_i \sim H_0(X_i)$$

$$u_i \sim F_0(u_i; \theta)$$

$$Y_i = y_0(X_i, u_i; \theta)$$

The data is (Y_i, X_i) with u_i unobserved.

We know this model up to θ

To think of this as non-parametric we can think of θ as infinite dimensional

For example if F_0 is nonparametric we could write the model as $\theta = (\theta_1, F_0(\cdot))$

Point Identification of the Model

The model is identified if there is a unique θ that could have generated the population distribution of the observable data (X_i, Y_i)

A bit more formally, let Θ be the parameter space of θ and let θ_0 be the true value

- If there is some other $\theta_1 \in \Theta$ with $\theta_1 \neq \theta_0$ for which the joint distribution of (X_i, Y_i) when generated by θ_1 is identical to the joint distribution of (X_i, Y_i) when generated by θ_0 then θ is not identified
- If there is no such $\theta_1 \in \Theta$ then θ is (point) identified

Set Identification of the Model

Define Θ_I as the identified set.

I still want to think of there as being one true θ_0

Θ_I is the set of $\theta_1 \in \Theta$ for which the joint distribution of (X_i, Y_i) when generated by θ_1 is identical to the joint distribution of (X_i, Y_i) when generated by θ_0 .

So another way to think about point identification is the case in which

$$\Theta_I = \{\theta_0\}$$

Identification of a feature of a model

Suppose we are interested not in the full model but only a feature of the model: $\psi(\theta)$

The feature is identified if there is a unique value of it consistent with the observed data

More formally

$$\Psi_I \equiv \{\psi(\theta) : \theta \in \Theta_I\}$$

Most interesting cases occur when Θ_I is a large set but Ψ_I is a singleton

In practice $\psi(\theta)$ could be something complicated like a policy counterfactual in which we typically need to first get θ and then simulate $\psi(\theta)$

However, often it is much simpler and we can just write it as a known function of the data.

I can think of the Imbens/Angrist framework as a data generation process.

The data is (Z_i, T_i, Y_i)

Here Z_i is my exogenous X_i so $H_0(X_i)$ in that case is

$$Pr(Z_i = 1) = \rho$$

$$Pr(Z_i = 0) = 1 - \rho$$

They are not explicit about u_i but I can add it as

$$u_i = \begin{bmatrix} S_i \\ Y_{0i} \\ Y_{1i} \end{bmatrix}$$

and we can write this joint distribution as

$$Pr(S_i = s)$$

and $F_0(Y_{0i} | S_i = s)$ and $F_1(Y_{1i} | S_i = s)$ as the distributions of Y_{0i} and Y_{1i} respectively conditional on S_i .

Then finally $\Upsilon_i = (T_i, Y_i)$.

Once we have done this the y_0 is trivial

$$\begin{aligned} \Upsilon_i &= y_0(Z_i, u_i) \\ &= \begin{cases} (1, Y_{1i}) & S_i = \clubsuit \\ (0, Y_{0i}) & S_i = \heartsuit \\ (0, Y_{0i}) & S_i = \diamond, Z_i = 0 \\ (1, Y_{1i}) & S_i = \diamond, Z_i = 1 \end{cases} \end{aligned}$$

What can we hope to identify?

What we get from data is essentially:

- Dist of Z_i : ρ
- Dist of T_i conditional on Z_i . This is μ_{\clubsuit} when $Z_i = 0$ and $\mu_{\clubsuit} + \mu_{\diamond}$ when $Z_i = 1$, so the μ 's are identified
- Then
 - ① The distribution of Y_i conditional of $Z_i = 0, T_i = 0$, which is the distribution of Y_{0i} conditional on $S_i \in \{\heartsuit, \diamond\}$
 - ② The distribution of Y_i conditional of $Z_i = 0, T_i = 1$, which is the distribution of Y_{1i} conditional on $S_i \in \{\clubsuit\}$
 - ③ The distribution of Y_i conditional of $Z_i = 1, T_i = 0$, which is the distribution of Y_{0i} conditional on $S_i \in \{\heartsuit\}$
 - ④ The distribution of Y_i conditional of $Z_i = 1, T_i = 1$, which is the distribution of Y_{1i} conditional on $S_i \in \{\clubsuit, \diamond\}$

Notice that clearly the full model and the average treatment effect is not identified. The data is silent about

- the distribution of Y_{0i} for $S_i = \clubsuit$
- the distribution of Y_{1i} for $S_i = \heartsuit$

From that sense if we want to get a conditional treatment effect our only hope is the \diamond_s .

To see this is identified note that

- From 2 $E(Y_{1i} | S_i = \clubsuit)$ is identified
- From 4 we can identify

$$E(Y_i | Z_i = 1, T_i = 1) = \frac{\mu_{\clubsuit} E(Y_{1i} | S_i = \clubsuit) + \mu_{\diamond} E(Y_{1i} | S_i = \diamond)}{\mu_{\clubsuit} + \mu_{\diamond}}$$

So

$$E(Y_{1i} | S_i = \diamond) = \frac{(\mu_{\clubsuit} + \mu_{\diamond}) E(Y_i | Z_i = 1, T_i = 1) - \mu_{\clubsuit} E(Y_{1i} | S_i = \clubsuit)}{\mu_{\diamond}}$$

is identified

- This is symmetric so

$$E(Y_{0i} | S_i = \diamond) = \frac{(\mu_{\heartsuit} + \mu_{\diamond}) E(Y_i | Z_i = 0, T_i = 0) - \mu_{\heartsuit} E(Y_{0i} | S_i = \heartsuit)}{\mu_{\diamond}}$$

- And thus

$$E(\alpha_i | S_i = \diamond) = E(Y_{1i} | S_i = \diamond) - E(Y_{0i} | S_i = \diamond)$$

is identified

Observations

- pretty nice that this is what IV actually converges to
- Not this simple if covariates or either Z_i or T_i is not binary
- however if Z_i take on lots of values you can get more, if it varies enough the average treatment effect is identified
- This doesn't work with other features of the distribution like the median. (difference of the medians is not the median of the difference)
- However mean is pretty general, we can define Y_i however we want so we can identify the mean of any function of Y_i .
 - One such function is $1(Y_i \leq y)$.
 - Do that at all y and we can identify the conditional cdf of Y_{1i} for the \diamond and the conditional cdf of Y_{0i} for the \diamond

